

Entry into Swedish Retail- and Wholesale Trade Markets

Sven-Olov Daunfeldt[¤], Niklas Rudholm^{¤¤} and Fredrik Bergström[¤]

[¤]The Swedish Research Institute of Trade (HUI)
S-103 29 Stockholm
Sweden

^{¤¤}Department of Economics
Umeå University
S-901 87 Umeå
Sweden

Abstract

In this paper, a Poisson regression model is used to study the determinants of entry of new firms into the Swedish retail- and wholesale trade markets. The results suggest that incumbent firms engage in strategic behavior to prevent the entry of newcomers. Regional factors are also found to be important determinants of entry.

Key words: Wholesale trade, retail trade, entry, number of firms, panel data.

JEL classification: L13, L81.

1 Introduction

This paper examines the entry process of retail and wholesale trade firms into the Swedish market between 1990 and 1996. It is generally believed (see e.g., Geroski, 1991) that new firms produce a number of benefits. For instance, the entry, or the threat of entry, of new firms is assumed to force prices down, thereby eliminating excess profits. Moreover, high entry rates may stimulate innovation, and increase productivity and product quality.

In Sweden, as in other countries, retail and wholesale trade are constantly changing sectors, and in recent years, for example, out-of-town shopping, chain stores' market shares, and the number of international competitors have all increased (see e.g., Bergström, 1999). However, the number of firms and total employment have increased more over the last decade in the wholesale than in the retail trade sector. Moreover, a relatively larger proportion of the employees in the wholesale trade sector work in a company owned by a firm headquartered outside Sweden (see Bergström et al., 2002).

Previous empirical studies (see e.g., Love, 1996; Audretsch and Fritsch, 1994; Keeble and Walker, 1994; and Dunne et al., 1988) have generally used aggregated data on firm entry or have focused excessively on manufacturing firms. This is unfortunate because the entry process may differ between industrial sectors. For instance, Berglund and Brännäs (2001) have found that the determinants of entry differ between eight studied industrial sectors in Sweden. Moreover, Troske (1996) and Pakes and Ericson (1998) present results indicating that new non-manufacturing firms' grow to the size of the incumbents more quickly than do new manufacturing firms. The latter study is to our knowledge the only one that has examined the dynamics of retail trade firms, while we have found no study examining the entry pattern of wholesale trade firms. As the non-manufacturing sector of the economy grows, it has become increasingly important to gather information on the entry process of these firms as well.

In this paper, a Poisson regression model is used to study the entry process of retail and wholesale trade firms in Sweden. In contrast to previous empirical studies, who have focused on either firm (branch) specific data or region specific data, we are able to access firm-, branch- and region-specific data at the municipality level. We believe this to be important because the entry process of new firms may be influenced both by firm and branch specific factors, as well as regional characteristics. In addition, this is to our knowledge the first paper analyzing new firm entry into the retail and wholesale trade

markets.

The results indicate that few new firms enter markets where incumbent firms face low profits and high sunk costs, or where the market has previously been characterized by the presence of a local monopoly. This is to be expected since economic theory predicts entry to be more common when the minimum efficient scale, i.e., sunk costs, are low. Moreover, monopoly firms can use excess profits to discourage entry in a number of ways. One way would be to use the excess profits for marketing efforts, thereby increasing brand loyalty. Our results suggests that incumbent firms may engage in strategic behavior in order to prevent the entry of new competitors. In addition, we found a number of regional variables that affect entry behavior in the retail and wholesale trade sectors.

The paper is organized as follows: the next section presents the theoretical framework of our study; section three describes the data used; section four presents the empirical model and the projected results; and section five presents the main conclusions and suggestions for future work.

2 The Model

We assume that firms enter a particular market with the intention of making a profit, and that potential entrants face entry costs. Bain (1956) defined barriers to entry as the extent to which established firms can elevate their selling price above their minimum average cost without stimulating entry into the industry. Some possible barriers to entry mentioned in the literature (see e.g., Spence, 1977; Dixit, 1979) include economies of scale or scope, product differentiation and absolute cost advantages of established firms.

Following Rudholm (2001), we assume that the markets are characterized by imperfect competition, and that firms enter a given market until the profit in each period is driven to zero, that is until

$$\pi_{it} = \sum_{k=1}^K p_{ikt}(Q_{kt})q_{ikt} - \sum_{k=1}^K C_{ikt}(q_{ikt}) - F_{it} = 0; \quad (1)$$

where π_{it} denotes the profit of the potential entrant i ($i = 1, \dots, n$) in period t ($t = 1990, \dots, 1996$). In equation (1), $p_{ikt}(Q_{kt})$ is the price of product k as a function of the total market sales of that product, q_{ikt} is the sales of product k of the potential entrant conditional on entry, and $C_{ikt}(q_{ikt})$ is the total sales

costs as a function of the sales volume of product k . Thus, $\frac{1}{2}F_{it}$ represents the total profit for a firm in the wholesale- or retailing trade business selling a total of K different products. The term F_{it} can be interpreted as the entry cost corresponding to the zero profit condition, i.e., when additional entrants are unable to make profits. This means that the profits of already established firms can be positive without attracting the entry of new firms. Let F_{it} take the following form:

$$F_{it} = c_0 + c_1T + \gamma_1\Phi N_t + \frac{1}{2}G_t + \beta^0 X_{jt} + \alpha^0 Y_{mt} + \varepsilon_{it} \quad (2)$$

where c is a constant term, T is a trend variable, and $\Phi N_t = N_t - N_{t-1}$ is a measure of net entry in period t , meaning that the entry cost is allowed to depend on recent entry by other firms. The entry cost also depends on a vector of general components (e.g., business cycles and institutional changes), G_t ; a vector reflecting branch specific explanatory variables, X_{jt} , and a vector reflecting the characteristics of the municipality where the firm is planning to set up operations, Y_{mt} ; where the subscripts j and m denote branch and municipality, respectively. Finally, the entry cost also contains a component, ε_{it} , not observed by the researcher. ε_{it} is interpreted as a realization from a distribution of a stochastic variable with zero mean and constant variance.

Define $\frac{1}{2}F_{it}^0 = \sum_{k=1}^K p_{ikt}(Q_{kt})q_{ikt} - \sum_{k=1}^K C_{ikt}(q_{ikt})$ as the profit opportunity of the potential entrant in the absence of the entry cost. Substituting equation (2) into equation (1) and solving for $\Phi N_t = N_t - N_{t-1}$, our measure of net entry, gives the following expression:

$$\Phi N_t = \theta_0 + \theta_1T + \gamma^0 \frac{1}{2}F_{it}^0 + \mu^0 G_t + \beta^0 X_{jt} + \alpha^0 Y_{mt} + \varepsilon_{it} \quad (3)$$

where $\theta_0 = \gamma_1 c_0 = \gamma_1'$; $\theta_1 = \gamma_1 c_1 = \gamma_1'$; $\gamma^0 = 1 - \gamma_1$; $\mu^0 = \gamma_1 (1 - \gamma_1)^{\frac{1}{2}}$; $\beta^0 = \gamma_1 (1 - \gamma_1)^{\frac{1}{2}}$ and $\alpha^0 = \gamma_1 (1 - \gamma_1)^{\frac{1}{2}}$. The parameter θ_0 is a constant term, γ^0 measures how potential profits affect entry, μ^0 captures general factors at the national level influencing firms' entry behavior, β^0 is a parameter vector corresponding to the branch-specific explanatory variables, and α^0 is a parameter vector corresponding to region-specific effects. Finally, $\varepsilon_{it} = \gamma_1 \varepsilon_{it} = \gamma_1'$ is a random disturbance term which, from the assumptions regarding ε_{it} , has zero mean and constant variance. The variables used in the estimation of equation (3), and thus related to equations (1) and (2) above, will be discussed thoroughly in the empirical section.

3 The Empirical Analysis

3.1 Data

We are able to access firm- and region-specific data at the municipality level. All Swedish firms are legally bound to submit their annual reports to the Swedish patent and registration office (PRV). Data from the annual reports of firms active in the wholesale and retail markets between 1989 and 1996 were used in this study. This data set was collected by Upplysningscentralen AB (UC).¹ The data include, among other items, measures of profits, salaries, fixed costs and liquidity.

The municipality-specific data were provided by Statistics Sweden. These data cover matters such as the tax levels, expenditures, government grants, demographics, average income, political preferences, educational level and unemployment in each municipality. Due to the division of some municipalities into smaller units, as well as the mergers of three counties in Sweden during the studied period, 56 municipalities were omitted from this study, leaving a total of 233.

The sample was restricted to firms with documented positive sales during the study period. Before aggregating the data to the proper branch level, our data set contained 31,448 retail trade firms, and 35,550 wholesale trade firms (1996). Aggregating the data to the five digit SNI-code level², we have access to a total of 44,791 observations pertaining to the retail trade sector, and 37,916 observations pertaining to the wholesale trade sector during the study period. In addition, 10,700 firms entered the retail trade sector during the study period, while 3,550 firms exited the sector. In the wholesale trade sector, there were 12,132 new entrants in the study period, while 3,690 firms exited the sector. Hence, we can observe a clear positive trend in the number of retail and wholesale trade firms during the study period.

3.2 Econometric Methods

Turning to the econometric specification of the model, two aspects of the data must be considered. First, since the number of firms entering a market is a

¹UC is a Swedish credit information firm that collects economic information on both firms and individuals residing in Sweden.

²SNI refers to the Swedish standard industrial classification.

positive integer, a count data model is used. Second, there are only a few non-zero observations on entry, that is, n_{it} is in many cases equal to n_{it-1} . The Poisson regression model meets these requirements and corresponds directly to the model presented in Section 2. Thus, we estimate a Poisson regression model where net entry is regressed on a number of independent variables separately for the retail and the wholesale trade sectors. As the focus of this paper is entry into a specific branch of business, all firm-specific data have been aggregated to the proper branch level j .³ More formally, the following model is estimated:

$$\text{Prob}(\Phi N_{jt}) = \frac{e^{i_{jt}} \Phi N_{jt}}{\Phi N_{jt}!}; \quad \Phi N_{jt} = 0; 1; 2; 3; \dots \quad (4)$$

where

$$\ln i_{jt} = \beta_0 + \beta_1 T + \beta_{j,t-1}^0 + \beta_j^1 F_{jt} + \mu_k^0 G_t + \beta_s^0 X_{jt} + \beta_z^0 Y_{mt} + u_{it} \quad (5)$$

Profit opportunities for the entrant are captured by $\beta_{j,t-1}^0$; which measures operating profits. Equation (5) also includes branch-specific fixed effects (F_{jt}) in the sense that different intercepts are estimated for each type of business. General factors at the national level that influence the entry behavior of individual firms are captured by G_t , X_{jt} contain characteristics of the incumbents that are assumed to prevent entry, Y_{mt} is a vector of regional determinants of entry, β_0 is a constant, u_{it} is a disturbance term with zero mean and constant variance, and β_1 , β_j , μ_g^0 ($g = 1; 2; 3$), β_s^0 ($s = 1; 2$), and β_z^0 ($z = 1; \dots; 13$) are parameters to be estimated.

The variable representing profit opportunities for entrants, as well as all branch-specific variables, have been lagged one period. Lagging these variables has two advantages. First, this corresponds directly to the potential entrant's decision problem, since entrants only have access to other firms' annual reports with a one year time lag. Second, this setup makes it possible to reduce a possible endogeneity problem. The data contains several profit measures. We have chosen to use operating profits since this measure corresponds to the profit/loss from the main line of business of the firm. The branch-specific fixed effects have been included in the study, because barriers

³For the retail trade sector $j = 1; 2; \dots; 68$; for the wholesale trade sector $j = 1; 2; \dots; 56$.

to entry may vary systematically with unobserved factors such as, marketing efforts.⁴

General determinants of entry at the national level, G_t , include the gross domestic product (GDP), and two dummy variables reflecting the 1995 decision to increase the minimum capital necessary for starting up a limited company from SEK 50,000 to SEK 100,000. We expect to find both transitory and permanent effects following on this change. First, potential entrants may have chosen to enter the market before the increased capital requirement was enacted. This transitory effect is captured by a dummy variable taking the value one in the year 1994, i.e., the last year a new firm could be started at a cost of SEK 50,000. Another dummy variable takes the value one for the 1995-1997 period, reflecting a regime shift in the cost of starting up new businesses. As the cost increased, this variable is expected to have a negative effect on entry.⁵

Branch-specific factors, X_{jt} , that are assumed to characterize barriers to entry include sunk costs and a dummy variable reflecting regions with local monopolies in the preceding period. Large sunk costs are believed to reflect a commitment by incumbent firms to stay in the market, as these investments cannot be recouped if a firm has to leave the market. We used the level of long-term fixed assets (in relation to total assets), such as buildings and machinery, as a proxy for the level of sunk costs. A dummy variable reflecting regions with a local monopoly in time period $t - 1$ was added because local monopolies can earn excess profits which later can be used to fund entry deterring behavior.

Net entry is, finally, also assumed to be determined by region-specific factors, Y_{mt} . The regional characteristics used in the estimation of equation (5) are population, population density, the municipality tax level, government investment grants, local government debt, average per capita income, the presence of a university or a university college, educational level of the population, net migration, political preferences, political stability and local unemployment.

Note that each municipality is formally free to set its tax level independently. However, during the recession years from 1991 to 1993 local gov-

⁴For an overview concerning how advertising can create an advantage for incumbents, see Comanor and Wilson (1979).

⁵A potential problem is that there might be other events affecting entry during these years. Hence, the parameter estimates for the dummy variables should be interpreted with caution.

ernments were temporarily deprived of this right. The availability of higher education is represented by a dummy variable assigned the value one if a university or a university college is located in the region. Data concerning educational level within the municipality refers to the percentage of the population that has at least enrolled in courses at a university or a university college. Political preferences are indicated by a dummy variable representing all local parliaments where non-socialist parties have the majority, while political strength is measured by the Herfindahl index.⁶

It is reasonable to expect the following results when estimating equation (5). First, intercept terms may differ among branches. This follows because the barriers to entry may vary systematically with unobserved factors, such as marketing efforts. From the theoretical framework, it follows that more entry should occur in branches where profits are high, while sunk costs and the presence of local monopolies should prevent entry.

Turning to the region-specific variables, population size, population density and average income in the municipality are used to measure the market size for the potential entrant. A number of previous studies (see e.g., Audretsch and Fritsch, 1994; Davidsson et al., 1994; and Guesnier, 1994) have indicated that more entry occurs in regions where markets are large. Audretsch and Fritsch (1994) among others have also found that entry is positively influenced by the level of education in the region, possibly indicating that firms demand highly skilled labor. More entry is, therefore, expected in municipalities with established universities and/or university colleges and where a high percentage of the population have enrolled in higher education. Entry may also be influenced by the local unemployment rate. Davidsson et al. (1994) note, however, that this effect may be positive or negative. A high level of unemployment may attract entry because this indicates that the firm has a large labor pool to draw from when recruiting. On the other hand, high unemployment may discourage entry because it serves as an indicator of lower regional demand in the future. Entry of new firms may also be influenced by the municipality tax rate. For instance, high taxes may prevent entry of new firms because a relatively large share of the local ser-

⁶The Herfindahl index is defined as

$$H = \sum_{p=1}^P \text{SHARE}_p^2$$

where SHARE is the percentage of representatives from party p.

vices are provided the public sector (see e.g., Fölster, 2000). On the other hand, Sinn (1996) argues that redistributive taxation promotes risk-taking and thereby stimulates self-employment. Turning to the variables concerning political preference, we expect entry to be more common in municipalities where there is a strong political leadership, as measured by the Herfindahl index, because this may create a stable working environment for the firm. In addition, firms may prefer a non-socialist local government because this leadership is likely to implement more beneficial policies for the firm compared to a socialistic local government. Hence, the type of the political leadership, socialist or non-socialist, might have an effect on entry.

3.3 Estimation Results

Equation (5) is estimated by means of an iterative maximum likelihood technique. The estimation results are presented in Table 1.

The estimate of the time trend is negative and significant for both sectors. Turning to the parameter estimate for our GDP index, these are negative and significant at the one-percent significance level for both sectors. This result is rather unexpected, but might reflect the fact that self-employment increased during the recession of the early 1990s.

The dummy variable reflecting the regime shift is positive and significant for both sectors, contrary to our expectations. Note, however, that the dummy variables reflecting the costs of starting new companies also reflect all other period-specific events that affect entry.

Turning to branch-specific variables, we find that there is more entry into branches where profits are high. The parameter estimates for profits are significant at the ten-percent level for retail trade, and at the one-percent level for wholesale trade businesses. Although this is as one would expect from microeconomic theory, it has not been widely reported in previous empirical studies of entry behavior (see e.g., Geroski, 1995). In this paper, we used the level of long-term fixed assets (in relation to total assets) such as buildings and machinery as a proxy for the level of sunk costs. The results indicate that higher sunk costs decrease entry into the retail trade sector. Finally, the parameter estimates regarding local monopolies are negative and statistically significant at the one-percent level for both sectors, indicating that entry is less likely in branches where the incumbent firm had a local monopoly in the previous period.

Table 1. Estimation results^a

Variable (parameter)	Retail		Wholesale	
	Estimate	t-value	Estimate	t-value
Constant (θ_0)	2.88	3.25	1.81	1.84
Trend (θ_1)	-0.11	-6.34	-0.14	-6.96
Pro...ts (θ)	2.69Ei ⁹	1.66	6.49Ei ¹⁰	4.74
GDP (μ_1)	-0.038	-4.29	-0.049	-4.91
Dcapital (μ_2)	0.23	3.47	0.21	2.87
Dcapital94 (μ_3)	0.021	0.35	-0.0037	-0.05
Sunk cost (γ_1)	-0.18	-2.39	-0.027	-1.60
Dmonopoly (γ_2)	-0.71	-22.71	-0.73	-23.17
Population (\pm_1)	4.9Ei ⁶	18.53	4.59Ei ⁶	19.45
Population density (\pm_2)	-0.00013	-3.20	0.000096	3.08
Migration (\pm_3)	-2.4Ei ⁶	-0.12	-0.000061	-2.81
Unemployment (\pm_4)	-0.000016	-2.84	-0.000011	-1.88
Education (\pm_5)	3.07	11.35	4.40	16.80
Income (\pm_6)	0.000024	1.03	0.000037	2.53
Tax rate (\pm_7)	0.018	3.36	0.031	5.37
Debt (\pm_8)	0.22	3.85	0.38	6.45
Investment grants (\pm_9)	-0.00019	-0.89	-0.000042	-0.15
Dconservative (\pm_{10})	0.024	0.82	0.13	3.81
Political strength (\pm_{11})	0.79	3.01	1.16	4.00
Duniversity (\pm_{12})	0.23	5.91	0.13	3.33
Number of obs	44791		37916	
Pseudo R ²	0.20		0.29	

*t-values are robust against overdispersion.

Turning to the variables reflecting regional differences, we find that entry is more common in areas with large populations. This result is statistically significant at the one-percent level for both sectors. However, the population density seem to have a negative effect on entry into retail trade, while it has the expected positive effect on entry into the wholesale trade sector.

Local unemployment rates have a negative impact on entry, suggesting

that unemployment may serve as an indicator of regional demand. The proportion of the population having a university and/or university college education has a positive impact on entry, and firms are more likely to enter markets where they have access to a large stock of well-educated workers. The variable reflecting average income per capita is positive for both sectors, but statistically significant only in the model concerning entry into wholesale trade markets.

The parameter estimate concerning local tax rates shows that entry is more common in regional markets where the tax rate is high. A variable we consider to be closely related to taxes is the level of debt in relation to total assets in the municipality. Municipalities with a large level of debt might be forced to raise taxes in the future to repay their loans. Firms considering entering a market in a particular area recognize this, and are thus reluctant to enter. However, we find that the parameter estimates are positive for both sectors, as well as statistically significant at the one-percent level. We have not been able to find any reasonable explanation for this unexpected result.

Turning to the variables reflecting political preference and the political strength of local government, we find that entry into wholesale trade markets is more common in municipalities governed by non-socialist parties. In addition, the parameter estimates show that entry is more common for both sectors in municipalities where there is a strong political leadership, irrespective of the ideological inclination of the government. Our results indicate, moreover, that municipalities with a university or a university college attract more entries. Note, finally, that the determinants of firm entry into a local market are similar for the retail and wholesale trade sectors.

4 Conclusions

This paper examines the determinants of firm entry into the Swedish retail and wholesale trade markets. The results suggest that entry is more common in markets where the incumbents have high profits, while entry is found to be less common in markets characterized by a single incumbent (i.e., markets subject to local monopolies). Moreover, high required investments in sunk costs were also found to discourage entry in the retail trade sector.

Regional factors are also found to influence entry behavior. According to the results, regions characterized by large populations, high educational level, the presence of a university college or university, large debts, high taxes and

strong political leadership are more likely to attract entry from other regions.

We observe a clear positive trend in the number of firms despite the recession during the years under study. This may indicate that firms have some other mechanism, aside from exiting, by which to adapt to changing market conditions. One such mechanism could, for instance, be investments/disinvestments in capital and/or labor. This is, therefore, an interesting question for further research.

5 Acknowledgments

Thomas Aronsson and the participants in a seminar held at Umeå University are gratefully acknowledged for their valuable comments. The first and third author acknowledges financial support from the Foundation for Research in Trade and Commerce. The second author thanks the Browaldh-Wallander-Hedelius foundation for their generous financial support.

References

- Audretsch, D.B., and Fritsch, M. (1994), "The Geography of Firm Births in Germany", *Regional Studies*, 28, 359-365.
- Bain, J. (1956) *Barriers to new competition*. Cambridge, Harvard University Press.
- Berglund, E., and Brännäs, K. (2001), "Plants' Entry and Exit in Swedish Municipalities", *The Annals of Regional Studies*, 35, 431-448.
- Bergström, F. (1999), "Does Out-of-Town Shopping Really Crowd Out High Street Shopping", working paper, the Swedish Research Institute of Trade (HUI).
- Bergström, F., Rämme, U., and Wengström, E. (2002), "Struktur och strukturmöjigheter i partihandeln", working paper, the Swedish Research Institute of Trade (HUI).
- Comanor, W.S., and Wilson, T.A. (1979), "Advertising and Competition: A Survey", *Journal of Economic Literature*, 17, 453-476.
- Davidsson, P., Lindmark, L., and Olofsson, C. (1994), "New Firm Formation and Regional Development in Sweden", *Regional Studies*, 28, 395-410.

- Dixit, A. (1979), "A Model of Duopoly: Suggesting a Theory of Entry Barriers", *Bell Journal of Economics*, 10, 20-32.
- Dunne, T., Roberts, M.J., and Samuelson, L. (1988), "Patterns of Firm Entry and Exit in U.S. Manufacturing Industries", *RAND Journal of Economics*, 19, 495-515.
- Fölster, S. (2000), "Do Lower Taxes Stimulate Self-employment?", working paper, the Swedish Research Institute of Trade (HUI).
- Geroski, P.A. (1991), *Market Dynamics and Entry*, Basil Blackwell, Oxford.
- Geroski, P.A. (1995), "What do we Know about Entry?", *International Journal of Industrial Organization*, 13, 421-440.
- Guesnier, B. (1994), "Regional Variations in New Firm Formation in France", *Regional Studies*, 28, 347-358.
- Keeble, D., and Walker, S. (1994), "New Firms, Small Firms and Dead Firms: Spatial Patterns and Determinants in the United Kingdom", *Regional Studies*, 28, 411-427.
- Love, J.H. (1996), "Entry and Exit: A County-level Analysis", *Applied Economics*, 28, 441-451.
- Pakes, A., and Ericson, R. (1998), "Empirical Implications of Alternative Models of Firm Dynamics", *Journal of Economic Theory*, 79, 1-45.
- Rudholm, N. (2001) "Entry and the Number of Firms in the Swedish Pharmaceuticals Market, *Review of Industrial Organization*, 19, 351-364.
- Sinn, H.W. (1996), "Social Insurance, Incentives and Risk-taking", *International Tax and Public Finance*, 3, 259-280.
- Spence, M. (1977), "Entry Capacity, Investment and Oligopoly Pricing", *Bell Journal of Economics*, 8, 534-544.
- Troske, K.R. (1996), "The Dynamic Adjustment Process of Firm Entry and Exit in Manufacturing and Finance, Insurance and Real Estate", *Journal of Law and Economics*, 39, 705-735.